# Online Statistical Inference in Decision-Making with Matrix Context



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### Objective in Sequential Decision-Making

- Regret Minimization
  - Bandit algorithms are designed for regret minimization: how much worse it performs compared to an offline oracle





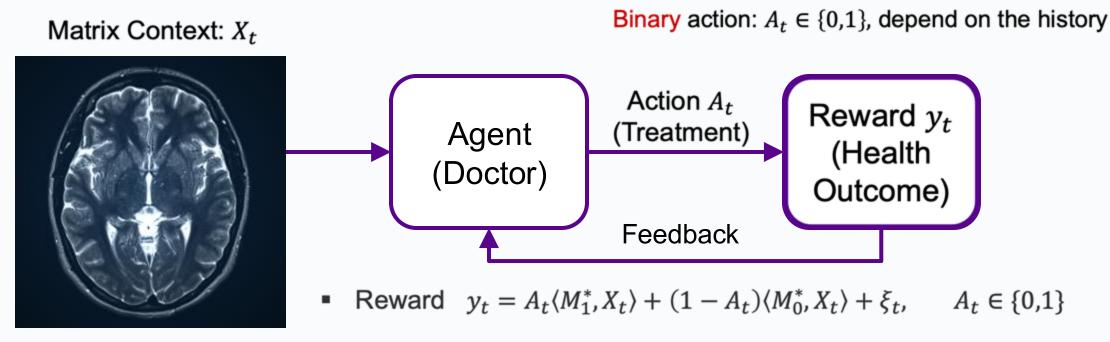




- In some cases, we also care about statistical inference or uncertainty quantification
  - Infer the effect of an arm / difference of effect between two bandit arms
    - Under-performing arms could be eliminated or modified, High-performing arms could be studied further, and One arm might be better than another
  - Gaining generalizable knowledge; Identifying new directions
  - Crucial for scientific discovery

### **Decision-Making with Matrix Context**

Some applications require the context information stored in a matrix



- $\widehat{M}_{i,t-1}$  is the estimated parameter for selecting the action
- $\varepsilon$ -greedy policy, for  $\varepsilon \in (0,1)$ , exploration-exploitation tradeoff.

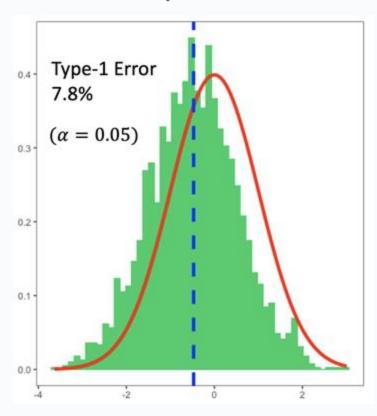
$$A_t \sim Ber(\pi_t), \quad \pi_t = (1 - \varepsilon)I\{\langle \widehat{M}_{1,t-1}, X_t \rangle > \langle \widehat{M}_{0,t-1}, X_t \rangle\} + \frac{\varepsilon}{2}$$

### Inference - Bandit Algorithms Induce Dependence

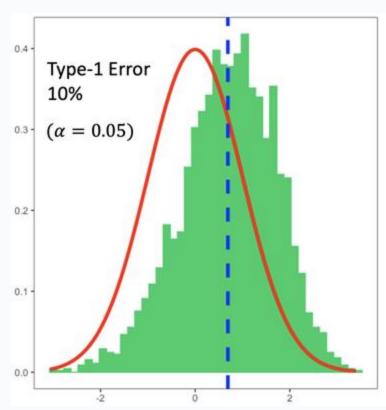
- Observations  $(X_t, A_t, y_t)$  are not independent over  $t \in \{1, 2, ..., n\}$ 
  - We use past history to select the action A<sub>t</sub> in the new context X<sub>t</sub>
  - Bandit data is adaptively collected
- Consequences for Statistical Inference e.g., test whether (1,2) -entry of matrix  $M_1$  is 0.  $H_0: M_1^*(1,2) = 0$ 
  - Violates independence assumptions of standard inference
  - Introduce a bias, or potentially non-normality
  - Villar, Bowden, Wason (2015, Stat. Sci.); Deshpande et al (2021, JASA); Khamaru, Mackey,
     Wainwright (2021); Zhang, Janson, Murphy (2021, 2022); Chen, Lu, Song (2021ab, JASA)
    - Not for matrix context, not fully-online algorithms.

### Illustration and Comparison of Different Sources of Bias

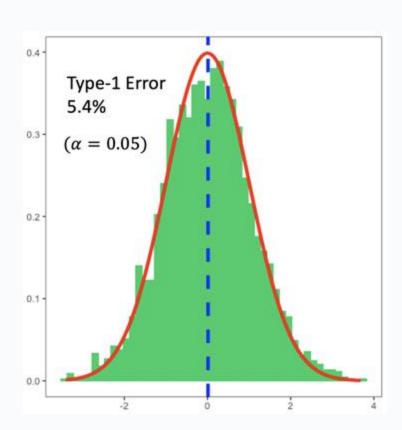
• In our problem, Bias in  $\widehat{M}$  comes from two sources:



Bias caused by adaptivity of data collection.



Bias caused by nonconvex optimization.



Our doubly-debiased estimator handling both sources of bias.

### Can we design estimators for $M_1^*$ , $M_0^*$ that are simultaneously:

sample-efficient, computation-efficient, & unbiased?

Low-rank matrix

Fully-Online

Inference

- Sequential/Fully online.
- Non-convexity.
- Noisy samples.
- Desire sample-efficient convergence rate,  $O_P(\sqrt{dr/t})$

- Fast computation
- Minimal storage

- Doubly-debiasing.
- Traditional analysis not applicable
- Consistent variance estimator

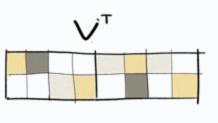
## Low-rank Matrix (Trace) Regression for Each Arm

- Arm i: Reward  $y_t = \langle M_i^*, X_t \rangle + \xi_t$
- Minimize the squared loss

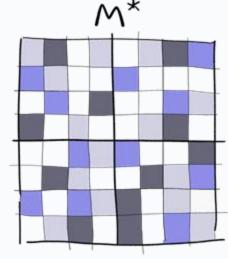
$$\min_{\mathcal{U}_i \in \mathbb{R}^{d_1 \times r}, \mathcal{V}_i \in \mathbb{R}^{d_2 \times r}} \frac{1}{2} \sum_{t=1}^n (y_t - \langle \mathcal{U}_i \mathcal{V}_i^\top, X_t \rangle)^2$$

Matrix Regression: Population Version





$$\boldsymbol{M_i} = \boldsymbol{U_i}\boldsymbol{V_i^{\top}}$$



$$\min_{\mathcal{U}_i \in \mathbb{R}^{d_1 \times r}, \mathcal{V}_i \in \mathbb{R}^{d_2 \times r}} \mathbb{E}_{(X,y) \sim \mathsf{P}_{X,y}} f(\mathcal{U}_i, \mathcal{V}_i; (X,y))$$

$$\min_{\mathcal{U}_i \in \mathbb{R}^{d_1 \times r}, \mathcal{V}_i \in \mathbb{R}^{d_2 \times r}} \mathbb{E}_{(X,y) \sim \mathsf{P}_{X,y}} \underline{f(\mathcal{U}_i, \mathcal{V}_i; (X,y))} \quad \begin{array}{l} \mathsf{Not \ convex \ in} \ (\underline{\mathit{U}_i, \mathit{V}_i})! \\ f(\mathcal{U}_i, \mathcal{V}_i; (X,y)) := \frac{1}{2} (y - \langle \mathcal{U}_i \mathcal{V}_i^\top, X \rangle)^2 \end{array}$$

Given a local initialization (e.g., Candès and Plan, 2010), an SGD Update for Arm 1:

$$\begin{pmatrix} \mathcal{U}_{i,t} \\ \mathcal{V}_{i,t} \end{pmatrix} = \begin{pmatrix} \mathcal{U}_{i,t-1} \\ \mathcal{V}_{i,t-1} \end{pmatrix} - \eta_t \begin{pmatrix} (\langle \mathcal{U}_{i,t-1} \mathcal{V}_{i,t-1}^\top, X_t \rangle - y_t) X_t \mathcal{V}_{i,t-1} \\ (\langle \mathcal{U}_{i,t-1} \mathcal{V}_{i,t-1}^\top, X_t \rangle - y_t) X_t^\top \mathcal{U}_{i,t-1} \end{pmatrix}. \qquad \eta_t = t^{-\alpha} \text{ step-size}$$

SGD Update: we only update  $U_i$ ,  $V_i$  for arm i with  $i = a_t$ 

### **Bandit SGD Update**

#### Renormalized SGD Update:

$$\tilde{\mathcal{U}}_{a_t,t-1} = W_{\mathcal{U}} D^{\frac{1}{2}} \text{ and } \tilde{\mathcal{V}}_{a_t,t-1} = W_{\mathcal{V}} D^{\frac{1}{2}}, \text{ where } W_{\mathcal{U}} D W_{\mathcal{V}}^{\top} \text{ is the top-}r \text{ SVD of } \mathcal{U}_{a_t,t-1} \mathcal{V}_{a_t,t-1}^{\top}.$$

$$\begin{pmatrix} \mathcal{U}_{i,t} \\ \mathcal{V}_{i,t} \end{pmatrix} = \begin{pmatrix} \tilde{\mathcal{U}}_{i,t-1} \\ \tilde{\mathcal{V}}_{i,t-1} \end{pmatrix} - \eta_t \frac{I\{a_t = i\}}{i\pi_t + (1-i)(1-\pi_t)} \nabla f(\tilde{\mathcal{U}}_{i,t-1}, \tilde{\mathcal{V}}_{i,t-1}; \{X_t, y_t\}).$$

#### Theorem 1

For any constant  $\alpha \in (0.5,1)$ , under some mild conditions and  $n = o(d^{\gamma})$  when  $d \to \infty$ , with high probability we have

$$\left\|\widehat{M}_{i,t}^{\mathsf{sgd}} - M_i^*\right\|_{\mathrm{F}} \leq C \gamma \sigma_i \sqrt{\frac{dr \log^2 d}{t^{\alpha}}}.$$

- Where  $d = \max\{d_1, d_2\}$  for  $d_1$  and  $d_2$  are the dimensions of  $M_i^*$ .
- r is the rank of  $M_i^*$  assumed known, and  $r \ll \min\{d_1, d_2\}$ .

for any  $1 \le t \le n$  and some positive constant C.

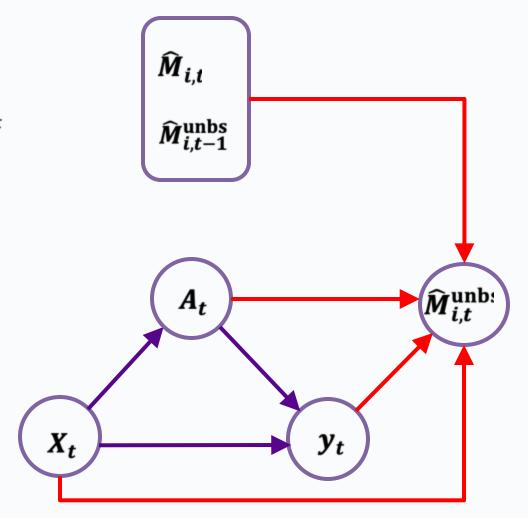
## Our Method: Step 2 (Sequential Debiasing)

• At time t, taking i = 1 for example, we debias

$$\widetilde{M}_{1,t} = \widehat{M}_{1,t-1} - \frac{I\{a_t = 1\}}{\pi_t} \left( \left\langle \widehat{M}_{1,t-1}, X_t \right\rangle - y_t \right) X_t$$
 Adaptive data collection Low-Rankness

- Running average:  $\widehat{M}_{1,t}^{\text{unbs}} = (\widetilde{M}_{1,t} + (t-1)\widehat{M}_{1,t-1}^{\text{unbs}})/t$
- Spectral Projection:

Project  $\widehat{M}_{i,t}^{\text{unbs}}$  to their leading singular subspaces (at the cost of negligible biases)



## **Estimator with Asymptotic Normality**

Parameter Inference

inference for the (1,2) entry of  $M_i$ 

$$\widehat{m}_{T}^{(i)} = \langle M_i, T \rangle$$
, e.g.  $T = e_1 e_2^{\mathsf{T}}$ 

Optimal Policy-Value Inference
 for optimal value V\* = E<sub>X</sub>[\langle M\_{\pi^\*(X)}, X \rangle]

#### Theorem 1: Asymptotic Normality (For i = 1)

Under mild conditions, with the presented decision-making model, as online sample size  $n \to \infty$ . Then

$$\sqrt{n} \Big( \widehat{m}_T^{(1)} - m_T^{(1)} \Big) \overset{d}{ o} \mathcal{N} \Big( 0, \sigma^2 S_1^2 \Big), ext{ where} \ S_1^2 = \int rac{\Big( \Big\langle U_ot U_ot^ op X V V^ op, T \Big
angle + \Big\langle U U^ op X V_ot V_ot^ op, T \Big
angle \Big)^2}{(1 - arepsilon_\infty) I \Big\{ \langle M_1 - M_0, X 
angle > 0 \Big\} + rac{arepsilon_\infty}{2}} dP_X.$$

### Theorem 2: Asymptotic Normality for $\widehat{V}_n$

Under mild conditions, as  $n \to \infty$ ,

$$\sqrt{n}\left(\widehat{V}_n-V^*
ight) \stackrel{d}{
ightarrow} \mathcal{N}(0,s^2),$$
 where

$$s^2 = \int \sum_{i=0}^1 \sigma_i I\{\langle M_i, X 
angle > \langle M_{1-i}, X 
angle\} dP_X + \mathbb{E}\left[\langle M_{\pi^*(X)}, X 
angle^2
ight].$$

### **Summary and Future Work**

- We propose a sample-efficient, computational efficient, & unbiased fully-online inference procedure for low-rank matrix contextual bandit with adaptive data collection.
- Two inference procedures for parameter inference, and optimal policy value inference.

 Qiyu Han, Will Wei Sun, and Yichen Zhang. Online Statistical Inference in Decision-Making with Matrix Context. Annals of Statistics, to appear. arXiv: 2212.11385.

